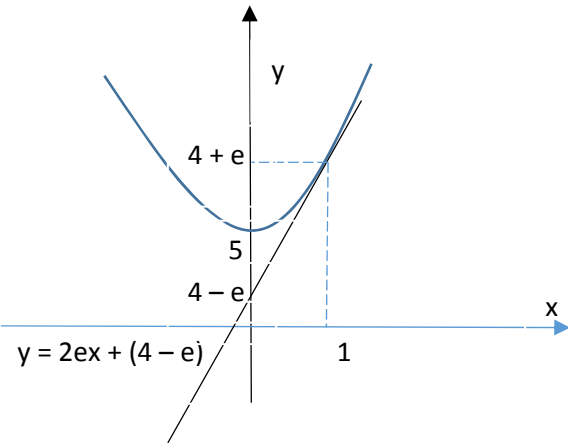
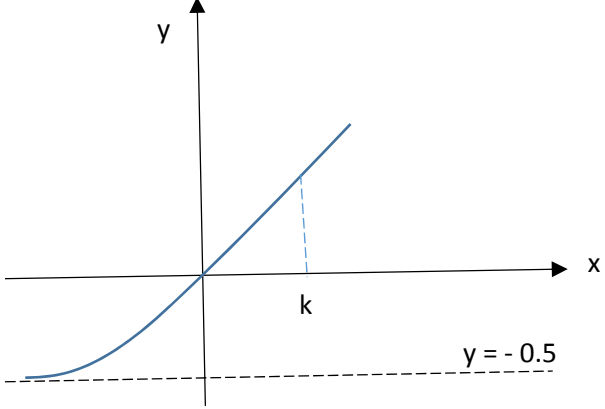




















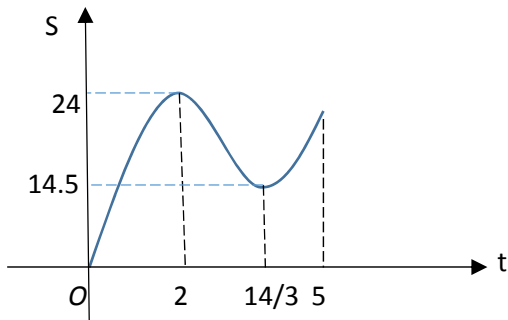
Solutions to 2017 Y6 H1 Maths Preliminary Exam II

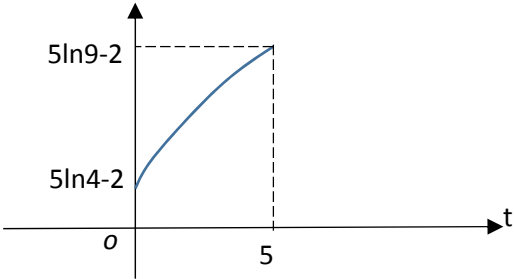
1	Working	
	$y = 4 + e^{x^2} \Rightarrow \frac{dy}{dx} = 2xe^{x^2}$ <p>when $x = 1, y = 4 + e, \frac{dy}{dx} = 2e$</p> $y - (4 + e) = 2e(x - 1)$ $y = 2ex + (4 - e)$  $\int_0^1 4 + e^{x^2} - (2ex + 4 - e) dx = 1.46$	

2	Working	
i)	$\frac{d \ln(x^2 + 1)}{dx} = \frac{1}{x^2 + 1} \cdot 2x = \frac{2x}{x^2 + 1}$	
ii)	$\frac{(x+1)^2}{x^2 + 1} = \frac{x^2 + 2x + 1}{x^2 + 1}$ $= \frac{x^2 + 1}{x^2 + 1} + \frac{2x}{x^2 + 1}$ $= 1 + \frac{2x}{x^2 + 1}$	
iii)	$\int \frac{(x+1)^2}{x^2 + 1} dx = \int 1 + \frac{2x}{x^2 + 1} dx = x + \ln(x^2 + 1) + c$	

3	Working	
i)	Let $x = a$ $\frac{1}{2}(e^a - 1) = a$ $e^a - 1 = 2a \Rightarrow e^a = 1 + 2a$	
ii)	 $\int_0^k \frac{1}{2}(e^x - 1) dx = \frac{1}{2} [e^x - x]_0^k$ $= \frac{1}{2}(e^k - k - 1)$ $\frac{1}{2}(e^k - k - 1) = \frac{a}{2}$ $e^k - k = 1 + a = e^a - a \text{ (from (i))}$ $\therefore k = a$	
iii)	Using GC, the points of intersection between $y = x$ and the curve is $x = 0$ and 1.256 The range of values of x : $x < 0$ or $x > 1.26$	

4	Working	
	<p>Let a, b and c be the interest rates of each fund in per cent.</p> $\frac{a}{100}(2000) + \frac{b}{100}(1500) + \frac{c}{100}(1000) = 309$ $20a + 15b + 10c = 309$ $2 \cdot \frac{a}{100}(2000) = \frac{b}{100}(1500) + \frac{c}{100}(1000)$ $40a - 15b - 10c = 0$ $\frac{a}{100}(2500) + \frac{b}{100}(2000) = 320$ $25a + 20b = 320$ <p>Solving the 3 equations, $a = 5.15$, $b = 9.56$ and $c = 6.26$ Fund B gave the highest rate = 9.56%</p>	

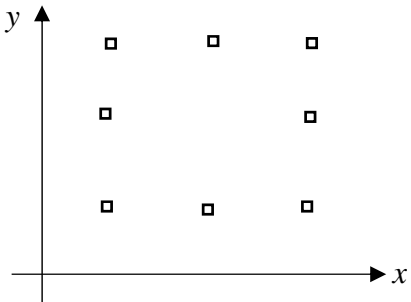
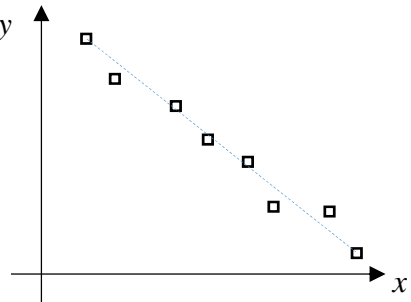
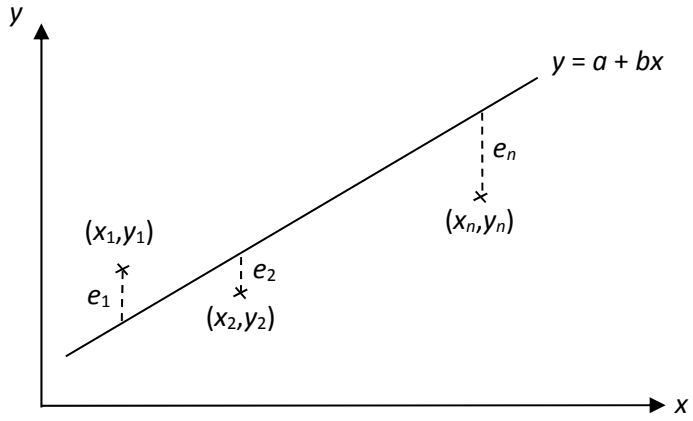
5	Working																									
i)	$\frac{dS}{dt} = 3t^2 - 20t + 28 = 0$ $t = \frac{20 \pm \sqrt{64}}{6} = 2 \text{ or } \frac{14}{3}$ <table><tr><td></td><td>1.995</td><td>2</td><td>2.005</td></tr><tr><td>gradient</td><td>+</td><td>0</td><td>-</td></tr><tr><td>shape</td><td></td><td></td><td></td></tr></table> <table><tr><td></td><td>4.665</td><td>14/3</td><td>4.675</td></tr><tr><td>gradient</td><td>-</td><td>0</td><td>+</td></tr><tr><td>shape</td><td></td><td></td><td></td></tr></table> <p>When $t = 2, S = 24$ which is maximum.</p>		1.995	2	2.005	gradient	+	0	-	shape					4.665	14/3	4.675	gradient	-	0	+	shape				
	1.995	2	2.005																							
gradient	+	0	-																							
shape																										
	4.665	14/3	4.675																							
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shape																										
ii)																										

iii)	$\int_0^5 t^3 - 10t^2 + 28t \, dt \approx 89.6$ <p>89.6 metres refers to the total distance covered by Johnson during 5 seconds.</p>	
iv)	$5 \ln(t+4) - 2 = 23$ $\ln(t+4) = 5$ $t+4 = e^5$ $t = e^5 - 4 \text{ seconds}$	
v)		
vi)	When $t = 1$, $\frac{d\sigma}{dt} = \frac{5}{1+4} = 1 \text{ cm}^3/\text{s}$	

6	Working	
6(i)	The probability = $\frac{1 \times 7 \times 1!}{9!} = \frac{1}{8 \times 9} = \frac{1}{72}$	
6(ii)	The probability = $\frac{(4 \times 3 \times 2!) \times 3!}{9!}$ $= \frac{1728}{362880} = \frac{1}{210}$	
6(iii)	The probability = $\frac{\binom{4}{2} \binom{3}{2} \binom{2}{1} \times 5!}{\binom{9}{5} \times 5!}$ $= \frac{36}{126} = \frac{2}{7}$	

	Working	Marks
7(i)	$P(X) = \frac{2}{6} = \frac{1}{3}$	
7(ii)	<p>The tree diagram for the events X and Y:</p> <pre> graph LR Root(()) --- 1/3 X((X)) Root --- 2/3 Xp1["X'"] X --- 3/4 Y1((Y)) X --- 1/4 Yp1["Y'"] Xp1 --- 5/12 Y2((Y)) Xp1 --- 7/12 Yp2["Y'"] </pre>	
7(iii)	$P(X Y) = \frac{P(X \cap Y)}{P(Y)}$ $= \frac{\frac{1}{3} \times \frac{3}{4}}{\frac{1}{3} \times \frac{3}{4} + \frac{2}{3} \times \frac{5}{12}}$ $= \frac{9}{19}$	
7(iv)	<p>Since $P(X Y) = \frac{9}{19}$, $P(X) = \frac{1}{3}$ and thus, $P(X Y) \neq P(X)$</p> <p>we can conclude that event A and B are not independent events.</p> <p>Alternative:</p> $P(Y) = \frac{1}{3} \times \frac{3}{4} + \frac{2}{3} \times \frac{5}{12} = \frac{19}{36}$ <p>Then $P(X \cap Y) = \frac{1}{4} \neq P(X) \times P(Y)$</p> <p>Hence, event A and B are not independent events.</p>	

8	Working	Marks																
8(i)	<p>The two assumptions are</p> <p>(i) The event of a particular chosen curry puff being too salty is independent of other curry puffs chosen by the MOH officials.</p> <p>(ii) The probability of choosing a curry puff that is too salty is constant for all curry puffs chosen by the MOH officials.</p> <p>(i) may not be true in general as curry puffs are manufactured in batches, thus if 1 curry puff is found to be too salty, curry puffs from the same batch will be salty too. (or any other similar reasoning)</p>																	
8(ii)	<p>Assume $X \sim B(n, 0.08)$,</p> <p>Then we need to have</p> $P(X = 1) < 0.01.$ <p>Using binomial pdf in GC, we can set the necessary and obtain the following probability table for checking:</p> <div><div><pre>Plot1 Plot2 Plot3 \Y1 Bbinompdf(X, 0 .08, 1) \Y2 = \Y3 = \Y4 = \Y5 = \Y6 =</pre></div><div><table><tr><th>X</th><th>Y1</th></tr><tr><td>42</td><td>.11006</td></tr><tr><td>43</td><td>.10367</td></tr><tr><td>44</td><td>.09759</td></tr><tr><td>45</td><td>.09183</td></tr><tr><td>46</td><td>.08636</td></tr><tr><td>47</td><td>.08118</td></tr><tr><td>48</td><td>.07627</td></tr></table><p>X=44</p></div></div> <p>Hence, from the above table, we deduce that the least value of n should be 44.</p>	X	Y1	42	.11006	43	.10367	44	.09759	45	.09183	46	.08636	47	.08118	48	.07627	
X	Y1																	
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8(iii)	<p>Let $n = 44$. Then $X \sim B(44, 0.08)$,</p> <p>Then we have</p> $P(3 \leq X < 10)$ $= P(X \leq 9) - P(X \leq 2).$ $= 0.692$																	

9	Working	Marks
9 (i)	<p>The required scatter diagram are as follows:</p> <p>(i) </p> <p>(ii) </p>	
9b(i)	<p>Let the sample of bivariate data be (x_i, y_i) where $i = 1, 2, \dots, n$.</p> <p>Let $e_i = y_i - (a + bx_i)$ be the vertical deviation between the point (x_i, y_i) and the line $y = a + bx$.</p> <p></p> <p>The line $y = a + bx$ is the least square regression line for the sample of bivariate data if the sum of the squares of the vertical deviations, i.e. $\sum_{i=1}^n (e_i)^2$, is the minimum.</p>	

	Using GC and the table of values, <div style="border: 1px solid black; padding: 5px; width: fit-content;"> <pre>LinReg y=a+bx a=.5590909091 b=.0774125874 r²=.9835387842 r=.991735239</pre> </div> <p>the regression line of y on x is $y = 0.559 + 0.0774x$</p>	
9b(ii)	a is the amount of cough syrup in ml produced with no amount of the chemical compound KH_2 added in the production process. b is the amount of cough syrup in ml produced with the addition of 1 mg of the chemical compound KH_2 added in the production process.	
9b(iii)	As x is the independent variable and y is the dependent variable in the data set, we will need to use the regression line of y on x to estimate the value of x given y value. Thus, when $y = 4.0$, $4.0 = 0.55909 + 0.077413x$ $\Rightarrow x = 44.448736 \approx 44.4$ mg (3 sig fig)	
9b(iv)	The estimate for x value in part 9b(iii) is reliable for the following reasons: i. the correlation coefficient $r = 0.992$ is very close to 1, ii. the given y value of 4.0 is within the data range of the y value iii. the correct line y on x is used as y is the dependent variable.	

10	Working	Marks
10(i)	The unbiased estimate for population mean $= \bar{x}$ $= \frac{720}{50} + 100 = 114.4$ The unbiased estimate for population variance $= s^2$ $= \frac{1}{50-1} \left(\sum (x-100)^2 - \frac{(\sum (x-100))^2}{50} \right)$ $= \frac{1}{49} \left(30500 - \frac{720^2}{50} \right) = \frac{2876}{7} = 410.8571429 \approx 411$ (3 sig. fig.)	
10(ii)	Let μ be the mean of the amount of energy released in each collision. For testing of the new claim by the young physicists, Let $H_0 : \mu = 108$ $H_1 : \mu \neq 108$ We then perform a one tail test at 5% level of significance i.e. $\alpha = 0.05$	

	<p>Under H_0, $\bar{X} \sim N\left(108, \frac{2876/7}{50}\right)$</p> <p>Then test statistics is $Z = \frac{\bar{X} - 108}{\sqrt{\frac{2876/7}{50}}} \sim N(0, 1)$</p> <p>Using GC, the p-value = 0.0255723191.</p> <p>Since p-value $< \alpha$, we reject H_0 and conclude that there is sufficient evidence at 5% level of significance that the population mean of the energy released in each collision is not 108 MeV.</p> <p>As the sample size n is large (50), it is not necessary to assume that X, the amount of energy released in each collision follows a normal distribution as by the Central Limit Theorem, \bar{X} can be approximated by a normal distribution for the test to be valid.</p>	
10(iii)	<p>For testing of the new claim by the young physicists, Let $H_0 : \mu = \mu_0$ $H_1 : \mu > \mu_0$</p> <p>We then need to perform a one tailed test at 3% level of significance i.e. $\alpha = 0.03$</p> <p>Under H_0, $\bar{X} \sim N\left(\mu_0, \frac{2876/7}{50}\right)$</p> <p>Then test statistics is $Z = \frac{\bar{X} - \mu_0}{\sqrt{\frac{2876/7}{50}}} \sim N(0, 1)$.</p> <p>For H_0 not to be rejected, $z_{\text{calculated}}$ must not be in the critical region, hence, $\frac{114.4 - \mu_0}{\sqrt{\frac{2876/7}{50}}} \leq 1.88079361$</p> $\Rightarrow 114.4 - \mu_0 \leq 1.88079361 \times \sqrt{\frac{2876/7}{50}}$ $\Rightarrow \mu_0 \geq 114.4 - 1.88079361 \times \sqrt{\frac{2876/7}{50}}$ $\Rightarrow \mu_0 \geq 109.0085999$ $\Rightarrow \mu_0 \geq 109 \text{ (3 sig fig)}$	

11	Working	Marks
i)	Let W be the r.v. of the amount of electricity used by a household in a month. $W \sim N(500, 80^2)$ $P(W < 600) = 0.89435 \approx 0.894$	
ii)	Let H be the number of households out of 10 which use less than 960 kWh, $H \sim B(10, 0.89435)$ $P(H \leq 8 H \geq 4) = \frac{P(4 \leq H \leq 8)}{P(H \geq 4)} = \frac{P(H \leq 8) - P(H \leq 3)}{1 - P(H \leq 3)}$ $= \frac{0.28583}{0.9999868} = 0.28583 \approx 0.286$	
iii)	Let X be the r.v. of number of months that use more than 600 kWh, $X \sim B(12, 0.10565)$. $P(X \geq 5) = 1 - P(X \leq 4) \approx 0.00550$	
iv)	$W_1 - W_2 \sim N(0, 12800)$ $P(W_1 - W_2 < 100) = P(-100 < W_1 - W_2 < 100) = 0.62324 \approx 0.623$	
v)	Assume that the amount of electricity used by a household of the two months is independent. This assumption may not be true if the family knows that they have over-used energy in the first month, they may use less in the later months. (or any other reasonable explanations)	
vi)	Let $C = 0.22W$ be the r.v. of the cost of electricity for a household in each month. $C \sim N(110, 309.76)$ $P(C > 150) = 0.011521 \approx 0.0115$	
vii)	$T = C_1 + C_2 + C_3 \sim N(330, 929.28)$ $P(T < m) \geq 0.96$ $m \geq 383.368 \approx 383$ (3 sig. fig)	
viii)	Let V be the r.v of the amount of electricity used by a household per month at the end of the campaign. Supposed $V \sim N(185, 80^2)$ Consider: a) $P(V < 0) = 0.0104$ Since the probability that a household uses negative amount of electricity is 0.0104 (1.04%) which is quite significant, Normal distribution is deemed as unsuitable. OR b) $P(185 - 3(80) < W < 185 + 3(80)) = 0.997$ Since probability of 0.997 of amount of electricity used is within 3 standard deviation and there is quite significant probability that the amount of electricity used is negative, Normal distribution is deemed unsuitable.	